

STA6934. PROBLEM 1.18

VLADIMIR BUGERA

Consider n observations x_1, \dots, x_n from \mathcal{B} where both k and p are unknown.

(a) Show that the maximum likelihood estimator of k , \hat{k} , satisfies

$$(1) \quad (\hat{k}(1 - \hat{p}))^n \geq \prod_{i=1}^n (\hat{k} - x_i)$$

and

$$(2) \quad ((\hat{k} + 1)(1 - \hat{p}))^n < \prod_{i=1}^n (\hat{k} + 1 - x_i),$$

where \hat{p} is the maximum likelihood estimator of p .

Density function of Binomial Distribution $\mathcal{B}(n, p)$ ($0 \leq p \leq 1$) is

$$(3) \quad f(x|p) = \binom{k}{x} p^x (1 - p)^{(k-x)}.$$

The likelihood function for the observations x_1, \dots, x_n is

$$(4) \quad L(k, p) = \prod_{i=1}^n f(x_i) = \prod_{i=1}^n \binom{k}{x_i} p^{x_i} (1 - p)^{(k-x_i)}$$

For maximum (\hat{k}, \hat{p}) of $L(k, p)$ the following should be valid:

$$(5) \quad L(\hat{k} - 1, \hat{p}) \leq L(\hat{k}, \hat{p}) \leq L(\hat{k} + 1, \hat{p})$$

Taking into account that

$$\begin{aligned} L(\hat{k} - 1, \hat{p}) &= \prod_{i=1}^n \frac{(\hat{k}-1)!}{x_i!(\hat{k}-1-x_i)!} \hat{p}^{x_i} (1 - \hat{p})^{(\hat{k}-1-x_i)}, \\ L(\hat{k}, \hat{p}) &= \prod_{i=1}^n \frac{\hat{k}!}{x_i!(\hat{k}-x_i)!} \hat{p}^{x_i} (1 - \hat{p})^{(\hat{k}-x_i)} \\ &= \prod_{i=1}^n \frac{\hat{k}}{(\hat{k}-x_i)} (1 - \hat{p}) \prod_{i=1}^n \frac{(\hat{k}-1)!}{x_i!(\hat{k}-1-x_i)!} \hat{p}^{x_i} (1 - \hat{p})^{(\hat{k}-1-x_i)}, \\ L(\hat{k}, \hat{p}) &= \prod_{i=1}^n \frac{(\hat{k}+1)!}{x_i!(\hat{k}+1-x_i)!} \hat{p}^{x_i} (1 - \hat{p})^{(\hat{k}+1-x_i)} = \\ &= \prod_{i=1}^n \frac{(\hat{k}+1)}{(\hat{k}+1-x_i)} (1 - \hat{p}) \prod_{i=1}^n \frac{(\hat{k})!}{x_i!(\hat{k}-x_i)!} \hat{p}^{x_i} (1 - \hat{p})^{(\hat{k}-x_i)}, \end{aligned}$$

we come to:

$$\begin{aligned} 1 &\leq \prod_{i=1}^n \frac{\hat{k}}{(\hat{k}-x_i)} (1 - \hat{p}), \\ \text{and} \\ 1 &\geq \prod_{i=1}^n \frac{(\hat{k}+1)}{(\hat{k}+1-x_i)} (1 - \hat{p}). \end{aligned}$$

After transformation we get:

$$\begin{aligned} (\hat{k}(1-\hat{p}))^n &\geq \prod_{i=1}^n (\hat{k} - x_i), \\ \text{and} \\ ((\hat{k}+1)(1-\hat{p}))^n &\leq \prod_{i=1}^n (\hat{k}+1 - x_i). \end{aligned}$$

Without losses of generality we can assume that the last inequality validates as strict. The equalities (1) and (2) are proved.

We will also need an expression connecting \hat{p} and \hat{k} . To find it we solve the following equation:

$$(6) \quad \frac{\partial \ln L(k, p)}{\partial p} \Big|_{p=\hat{p}, k=\hat{k}} = 0;$$

$$(7) \quad \frac{\partial \left(\ln \left(\prod_{i=1}^n \left(\frac{\hat{k}!}{x_i!(\hat{k}-x_i)!} \right) \right) + \ln \left(\prod_{i=1}^n \hat{p}^{x_i} (1-\hat{p})^{(\hat{k}-x_i)} \right) \right)}{\partial \hat{p}} = 0;$$

$$(8) \quad \sum_{i=1}^n \left(\frac{x_i}{\hat{p}} - \frac{(\hat{k}-x_i)}{1-\hat{p}} \right) = 0;$$

$$(9) \quad \frac{\sum_{i=1}^n x_i}{\hat{p}(1-\hat{p})} - \frac{n\hat{k}}{1-\hat{p}} = 0;$$

$$(10) \quad \hat{p} = \frac{\sum_{i=1}^n x_i}{\hat{k}n};$$

(b) If the sample is 16, 18, 22, 25, 27, show that $\hat{k}=99$.

Compute logarithms for both sides of equalities (1) and (2):

k	97	98	99	100	101	190
$\ln((k(1-\hat{p}))^n)$	21.643	21.694	21.745	21.795	21.845	25.004
$\ln(\prod_{i=1}^n (k - x_i))$	21.607	21.673	21.738	21.802	21.866	25.630
$\ln(((k+1)(1-\hat{p}))^n)$	21.694	21.745	21.795	21.845	21.894	25.031
$\ln(\prod_{i=1}^n (k+1 - x_i))$	21.673	21.738	21.802	21.866	21.928	25.660

Assuming that the likelihood function is unimodal we find $\hat{k} = 99$

(c) If the sample is 16, 18, 22, 25, 28, show that $\hat{k}=190$.

Compute logarithms for both sides of equalities (1) and (2):

k	99	188	189	190	191	192
$\ln((k(1-\hat{p}))^n)$	22.366	25.573	25.599	25.626	25.652	25.678
$\ln(\prod_{i=1}^n (k - x_i))$	21.724	25.564	25.594	25.624	25.654	25.683
$\ln(((k+1)(1-\hat{p}))^n)$	22.416	25.599	25.626	25.652	25.678	25.704
$\ln(\prod_{i=1}^n (k+1 - x_i))$	21.788	25.594	25.624	25.654	25.683	25.713

Assuming that the likelihood function is unimodal we find $\hat{k} = 190$

Conclusion: Maximum Likelihood Estimator is not robust in the presence of errors. The example demonstrates that a small deviation in data can increase an estimation more than two times. For more details the reader can be addressed to the following article:

Olkin, I., Petkau, A.J. and Zidek, J.V. (1981) A comparison of n estimators for the binomial distribution. J. Amer. Statist. Assoc., 76, 637-642