

Problem 7.29

a.) The transition from $x^* \rightarrow x$ is described by

$$K(x^*, x) = f_{X|X}(x|x^*) = \int_Y f_{Y|X}(y|x^*) f_{X|Y}(x|y) dy, \quad (1)$$

where $f_{Y|X}(y|x^*)$ is the density of $Y|X = x^*$, which is $N(\rho x^*, 1 - \rho^2)$ and $f_{X|Y}(x|y)$ is the density of $X|Y = y$, which is $N(\rho y, 1 - \rho^2)$. Plugging in the densities in (1) yields result (a).

b.) In general

$$\begin{aligned} & \int_{X^*} K(x^*, x) f_X(x^*) dx^* \\ &= \int_{X^*} \int_Y f_{Y|X}(y|x^*) f_{X|Y}(x|y) dy f_X(x^*) dx^* \\ &= \int_Y \int_{X^*} f_{Y|X}(y|x^*) f_X(x^*) dx^* f_{X|Y}(x|y) dy \\ &= \int_Y f_Y(y) f_{X|Y}(x|y) dy \\ &= f_X(x), \end{aligned}$$

hence the marginal distribution of X is the invariant distribution. In general, for bivariate normal (X, Y) , the marginal distribution of X is $N(\mu_X, \sigma_X)$, which in our case is the $N(0, 1)$ distribution. (See, for instance, Mood, Graybill and Boes (1974) for the derivative of the marginal distributions of the bivariate normal).

c.) We have to find the transition kernel $K(x^*, x) = f_{X|X}(x|x^*)$, i.e., solve the integral in (1). The exponents in the integral are, leaving aside the common factor $-\frac{1}{2(1-\rho^2)}$,

$$\begin{aligned}
& [(y - \rho x^*)^2 + (x - \rho y)^2] \\
= & [y^2(1 + \rho^2) - 2\rho y(x + x^*)] + [x^2 + \rho^2 x^{*2}] \\
= & (1 + \rho^2)[y^2 - 2y \frac{\rho}{1 + \rho^2}(x + x^*) + (\frac{\rho}{1 + \rho^2}(x + x^*))^2] \\
& - \frac{\rho^2}{1 + \rho^2}(x + x^*)^2 + [x^2 + \rho^2 x^{*2}] \\
= & (1 + \rho^2)[y - \frac{\rho}{1 + \rho^2}(x + x^*)]^2 + \frac{1}{1 + \rho^2}[-\rho^2(x + x^*)^2 \\
& + (1 + \rho^2)(x^2 + \rho^2 x^{*2})].
\end{aligned}$$

Now, think of $\frac{\rho}{1+\rho^2}(x + x^*)$ as a mean and $\frac{1-\rho^2}{1+\rho^2}$ as a variance, then the integral in (a) is one and we are left with the remaining constant factors, i.e.,

$$\begin{aligned}
f_{X|X}(x|x^*) = & - \frac{1}{2\pi(1 - \rho^2)} \sqrt{2\pi \frac{1 - \rho^2}{1 + \rho^2}} \exp\left\{-\frac{1}{2(1 - \rho^2)}\right. \\
& \left. \times \frac{1}{1 + \rho^2}[-\rho^2(x + x^*)^2 + (1 + \rho^2)(x^2 + \rho^2 x^{*2})]\right\},
\end{aligned}$$

which, after writing out the term in squared brackets and simplifying, results in the $N(\rho^2 x^*, 1 - \rho^4)$ density function.

d.) Since $X_k|X_{k-1} = x_{k-1} \sim N(\rho^2 x_{k-1}, 1 - \rho^4)$, X_k can be written in random walk form

$$X_k = \rho^2 X_{k-1} + U_k.$$

Now,

$$\begin{aligned} Cov(X_0, X_k) &= Cov(X_0, \rho^2 X_{k-1} + U_k) \\ &= Cov(X_0, \rho^2 X_{k-1}) + Cov(X_0, U_k) \\ &= \rho^2 Cov(X_0, X_{k-1}). \end{aligned}$$

Repeating the argument with k replaced by $k - 1$ gives

$$Cov(X_0, X_{k-1}) = \rho^2 Cov(X_0, X_{k-2}).$$

Hence, running through all k 's, finally

$$\begin{aligned} Cov(X_0, X_k) &= \rho^2 Cov(X_0, X_{k-1}) \\ &= \rho^2 \rho^2 Cov(X_0, X_{k-2}) \\ &= \dots \\ &= \rho^{2k} Var(X_0) = \rho^{2k}, \end{aligned}$$

which, as k goes to infinity, approaches 0, if $|\rho| < 1$.